

Supplementary Material

Simultaneous Bayesian Estimation of Excitatory and Inhibitory Synaptic Conductances by Exploiting Multiple Recorded Trials

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Appendix I. Problem Formulation

The dynamical system representing the dynamics of each recorded membrane potential (MP) is defined as follows:

$$\begin{cases} \mathbf{x}^i(t+1) = f(\mathbf{x}^i(t)) + \mathbf{v}(t) \\ y^i(t+1) = \hat{C}\mathbf{x}^i(t+1) + \sqrt{dt}\sigma_y \varepsilon^i(t) \end{cases}, \quad (\text{A.1})$$

where $\mathbf{v}(t)=[w(t), N_E(t), N_I(t)]^H$, index i corresponds to the i^{th} Kalman filter that is derived for the i^{th} trial, and

$$\begin{aligned} \mathbf{x}^i(t) &= [V^i(t), g_E^i(t), g_I^i(t)]^H \\ \mathbf{X} &= [\mathbf{x}^1(0:T), \dots, \mathbf{x}^i(0:T), \dots, \mathbf{x}^L(0:T)] \end{aligned} \quad (\text{A.2})$$

Both excitatory and inhibitory synaptic inputs (SIs) are approximated by a Gaussian distribution:

$$\begin{aligned} N_E(t) &\approx \text{Gaussian}(N_E(t); \boldsymbol{\mu}_{NE}(t), \boldsymbol{\Gamma}_{NE}(t)), \\ N_I(t) &\approx \text{Gaussian}(N_I(t); \boldsymbol{\mu}_{NI}(t), \boldsymbol{\Gamma}_{NI}(t)) \end{aligned} \quad (\text{A.3})$$

At each iteration, the algorithm runs L Kalman filters independently in order to estimate the $\mathbf{x}^i(1:T)$, for $i=1, \dots, L$. Given all the $\mathbf{x}^i(1:T)$, we aim to infer the mean and variance of the $N_E(1:T)$ and $N_I(1:T)$. The algorithm iterates using the estimated mean and variance of these SIs.

Appendix II. Inference of statistical parameters

We run Kalman filters (see details in Appendix 4 of [1]) for L trials to compute sufficient statistics required for EM algorithm.

We derive EM for our problem as described below.

$$\begin{aligned}
& E_{p(N_E, N_I | Y)} \left\{ \sum_i \log(p(y^i, X^i | \hat{\theta})) / Y, \theta \right\} = \\
& \sum_i \sum_{t=1}^T E_{p(N_E, N_I | Y)} \left\{ -\frac{1}{2} \log \sigma_\varepsilon^2 + (y^i(t) - V^i(t))^H (\sigma_\varepsilon^2)^{-1} (y^i(t) - V^i(t)) \right\} + \\
& \sum_i \sum_{t=1}^T E_{p(N_E, N_I | Y)} \left\{ -\frac{1}{2} \log \Gamma_{NE}(t) + (N_E^i(t) - \mu_{NE}(t))^H (\Gamma_{NE}(t))^{-1} (N_E^i(t) - \mu_{NE}(t)) \right\} + \\
& \sum_i \sum_{t=1}^T E_{p(N_E, N_I | Y)} \left\{ -\frac{1}{2} \log \Gamma_{NI}(t) + (N_I^i(t) - \mu_{NI}(t))^H (\Gamma_{NI}(t))^{-1} (N_I^i(t) - \mu_{NI}(t)) \right\}
\end{aligned} \tag{A.4}$$

Note that we aim to estimate a common SI for all given trials, where the dynamics of each trial produce their own synaptic perturbation (mean and variance of SIs underlying each single trial) as expressed below.

$$\begin{aligned}
N_E^i(t) &= \mu_{NE}^i(t) + \sqrt{\Gamma_{NE}^i(t)} \xi_E^i(t), \quad \xi_E^i(t) = \text{Gauss}(0,1) \\
N_I^i(t) &= \mu_{NI}^i(t) + \sqrt{\Gamma_{NI}^i(t)} \xi_I^i(t), \quad \xi_I^i(t) = \text{Gauss}(0,1)
\end{aligned} \tag{A.5}$$

The terms μ_{NE}^i , μ_{NI}^i , Γ_{NE}^i , and Γ_{NI}^i indicate the conditional mean (μ) and variance (Γ) of the SI given the observation of the i^{th} trial, namely, $\mu_{NE,I}^i = E\{N_{E,I}^i(t) | Y^i\}$ and $\Gamma_{NE,I}^i = \text{var}\{N_{E,I}^i(t) | Y^i\}$.

Note that we use (A.5) to simply express the conditional probabilities of SIs in (A.4).

By taking the derivative of (A.4) with respect to μ_N and Γ_N (for both excitatory and inhibitory inputs), the EM algorithm results in estimating the statistical parameters of the common SIs as follows:

$$\begin{aligned}
\mu_{NE}(t) &= \frac{1}{L} \sum_{i=1}^L \mu_{NE}^i(t) \\
\Gamma_{NE}(t) &= \frac{1}{L} \sum_{i=1}^L \{ \Gamma_{NE}^i(t) + (\mu_{NE}(t) - \mu_{NE}^i(t))^2 \}
\end{aligned} \tag{A.6}$$

$$\begin{aligned}
\mu_{NI}(t) &= \frac{1}{L} \sum_{i=1}^L \mu_{NI}^i(t) \\
\Gamma_{NI}(t) &= \frac{1}{L} \sum_{i=1}^L \{ \Gamma_{NI}^i(t) + (\mu_{NI}(t) - \mu_{NI}^i(t))^2 \}
\end{aligned} \tag{A.7}$$

To finish our EM section, we calculate, for excitatory and inhibitory SIs, the conditional mean of each trial, i.e., μ_{NE}^i and μ_{NI}^i , and the trial average of the conditional variances, i.e., $\frac{1}{L} \sum_{i=1}^L \Gamma_{NE}^i$

and $\frac{1}{L} \sum_{i=1}^L \Gamma^i_{NI}$. Note that all sufficient (conditional) statistics of SCs have been calculated previously by KF.

Consider the dynamics of SCs in (1):

$$g_{E,I}(t+1) = (1 - dt/\tau_{E,I})g_{E,I}(t) + N_{E,I}(t) \quad (\text{A.8})$$

The mean and variance of SCs can be calculated as follows (see [2] for more details):

$$\begin{aligned} E\{g_{E,I}(t)\} &= (1 - dt/\tau)E\{g_{E,I}(t-1)\} + E\{N_{E,I}(t-1)\} \\ \text{var}\{g_{E,I}(t)\} &= (1 - 2dt/\tau) \text{var}\{g_{E,I}(t-1)\} + \text{var}\{N(t-1)\} \end{aligned} \quad (\text{A.9})$$

Using the generic form of (A.9) for our problem, we have

$$\begin{aligned} \mu^i_{NE}(t) &= E\{g^i_E(t+1) | Y^i\} - (1 - dt/\tau_E)E\{g^i_E(t) | Y^i\} \\ \mu^i_{NI}(t) &= E\{g^i_I(t+1) | Y^i\} - (1 - dt/\tau_I)E\{g^i_I(t) | Y^i\} \\ \frac{1}{L} \sum_{i=1}^L \Gamma^i_{NE}(t) &\approx \frac{1}{L} \sum_{i=1}^L \left[\text{var}\{g^i_E(t+1) | Y^i\} - (1 - 2dt/\tau_E) \text{var}\{g^i_E(t) | Y^i\} \right] \\ \frac{1}{L} \sum_{i=1}^L \Gamma^i_{NI}(t) &\approx \frac{1}{L} \sum_{i=1}^L \left[\text{var}\{g^i_I(t+1) | Y^i\} - (1 - 2dt/\tau_I) \text{var}\{g^i_I(t) | Y^i\} \right] \end{aligned} \quad (\text{A.10})$$

The approximations in the previous two equations become more accurate if the number of trials L is large. Again, note that the statistics on the right-hand sides of these equations are calculated by KF.

References

- [1] Lankarany, M., Zhu, W. P., Swamy, M. N., and Toyoizumi, T. (2013). Inferring trial-to-trial excitatory and inhibitory synaptic inputs from membrane potential using Gaussian mixture Kalman filtering. *Front. Comput. Neurosci.* 7, 109. doi: 10.3389/fncom.2013.00109
- [2] Gillespie, D. T. (1996). The mathematics of Brownian Motion and Johnson noise. *Am. J. Phys.* 64, 225–240.